

## 47 8. Basis probabilistic concepts

### 8.1. Probability space

$(\Omega, \mathcal{A}, \mathbb{P})$  probability space

$\Omega$  sample space, space of outcomes  
 $\omega \in \Omega$  elementary outcome or event

$\mathcal{A}$   $\sigma$ -algebra or " $\sigma$ -field"

$A \in \mathcal{A}$  event

$\mathbb{P}$  probability measure defined  
on  $\mathcal{A}$ ,  $\mathbb{P} \geq 0$ ,  $\mathbb{P}(\Omega) = 1$ .

Ex. 8.2.  $\Omega = \{1, 2, \dots, 6\}$

$\mathcal{A} = \mathcal{P}(\Omega)$ ,  $\mathbb{P} = \frac{1}{6}$  counting measure

Pick  $\omega \in \Omega$  "at random"  
= roll a die

### 8.3. Random variables

A measurable map  $X: \Omega \rightarrow \mathbb{R}$   
is called a random variable

$(X^{-1}(B)) \in \mathcal{A}$  for each Borel set  
 $B \subseteq \mathbb{R}$

$$E[X] := \int_{\Omega} X(\omega) d\mathbb{P}(\omega)$$

(48) expectation of  $\bar{X}$ .

$$\begin{aligned}\text{Var}(\bar{X}) &= \mathbb{E}[(\bar{X} - \mathbb{E}[\bar{X}])^2] \\ &= \int (\bar{X} - \mathbb{E}[\bar{X}])^2 d\mathbb{P} \\ &= \mathbb{E}[\bar{X}^2] - \mathbb{E}[\bar{X}]^2.\end{aligned}$$

"Variance of  $\bar{X}$ "

Lev. 8.4. (Borel-Cantelli: I)

$A_n, n \in \mathbb{N}$ , events:

If  $\sum_{n=1}^{\infty} \mathbb{P}(A_n) < \infty$ , then

$$\mathbb{P}(A_n \text{ i.o.}) = 0$$

infinitely often

$= \{\omega \in \Omega : \omega \in A_n \text{ for inf. many } n\}$

$$= \bigcap_k \bigcup_{n \geq k} A_n$$

Proof:  $\mathbb{P}(A_n \text{ i.o.}) = \lim_{k \rightarrow \infty} \mathbb{P}\left(\bigcup_{n \geq k} A_n\right)$

$$\leq \limsup_{k \rightarrow \infty} \sum_{n=k}^{\infty} \mathbb{P}(A_n) = 0. \quad \square$$

Lev. 8.5. (Chebyshev's Ineq.)

If  $\bar{X} \geq 0$ , then

$$\mathbb{P}(\bar{X} \geq a) \leq \frac{\mathbb{E}[\bar{X}]}{a}, \quad a > 0.$$

49) Proof:

$$\mathbb{P}(X \geq a) = \int_{\Omega} \chi_{X \geq a}(\omega) d\mathbb{P}(\omega)$$

$$\leq \int_{\Omega} \frac{1}{a} X d\mathbb{P} = \frac{E[X]}{a} \quad \square$$

## 8.6. The distribution of a random variable

$X: \Omega \rightarrow \mathbb{R}^n$  random variable.

The distribution or law of  $X$  is the push-forward measure

$$\mathbb{P}_X := X_* \mathbb{P} \text{ on } \mathbb{R}^n, \text{ i.e.,}$$

$$\mathbb{P}_X(B) = \mathbb{P}(X^{-1}(B)) \text{ for each Borel set } B \subseteq \mathbb{R}^n.$$

$$E[X] = \int_{\mathbb{R}^n} x d\mathbb{P}_X(x)$$

Characteristic function of  $X: \Omega \rightarrow \mathbb{R}^n$

$$f(u) = E[e^{i(u \cdot X)}] \text{ for } u \in \mathbb{R}^n$$

$$f(u) = \int_{\Omega} e^{i(u \cdot X(\omega))} d\mathbb{P}(\omega)$$

$$= \int_{\mathbb{R}^n} e^{i(u \cdot v)} d\mathbb{P}_X(v)$$

= Fourier transform of its distribution.

(50)  $X_1, \dots, X_n: \Omega \rightarrow \mathbb{R}$  random variables.

$$\underline{X} := (X_1, \dots, X_n): \Omega \rightarrow \mathbb{R}^n.$$

Law of  $\underline{X} \stackrel{:=}{=} \frac{\text{joint law of } X_1, \dots, X_n.}{}$

### 8.7. Independence

$A, B \in \mathcal{A}$  events.

$A$  and  $B$  are independent if

$$\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

$$A^c := \Omega \setminus A.$$

If  $A, B$  ind., then  $A^c, B$  ind.:

$$\begin{aligned} \mathbb{P}(A^c \cap B) &= \mathbb{P}(B) - \mathbb{P}(A \cap B) \\ &= \mathbb{P}(B)(1 - \mathbb{P}(A)) = \mathbb{P}(A^c) \cdot \mathbb{P}(B) \end{aligned}$$

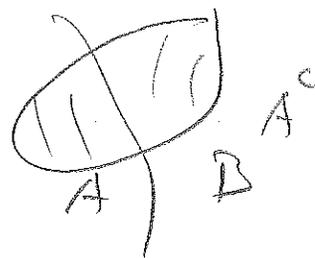
$\mathcal{F}_1, \dots, \mathcal{F}_n \subseteq \mathcal{A}$   $\sigma$ -algebras are independent if

$$\mathbb{P}(A_1 \cap \dots \cap A_n) = \mathbb{P}(A_1) \cdot \dots \cdot \mathbb{P}(A_n)$$

whenever  $A_i \in \mathcal{F}_i, i=1, \dots, n$ .

$A, B$  ind. iff  $\sigma$ -alg. gen. by  $A$  and  $B$  are independent.

$X_1, \dots, X_n$  random variables.



(51) They are independent  
 if the  $\sigma$ -algebras generated by  
 them are independent.

$$\sigma(\underline{X}) = \{ \underline{X}^{-1}(B) : B \in \mathbb{R}^h \text{ Borel} \}$$

If  $X_1, \dots, X_n$  are independent,  
 and  $f_1, \dots, f_n : \mathbb{R} \rightarrow \mathbb{R}$  are Borel,  
 then  $f_1(X_1), \dots, f_n(X_n)$  are  
 independent.

Note  $\sigma(f(\underline{X})) \subseteq \sigma(\underline{X})$ .

Thm. 8.8.  $X_1, \dots, X_n : \Omega \rightarrow \mathbb{R}$  random  
 variables,  $\underline{X} = (X_1, \dots, X_n) : \Omega \rightarrow \mathbb{R}^h$ .

TFAE:

(i)  $X_1, \dots, X_n$  are independent,

(ii)  $\mathbb{P}(X_1 \in B_1, \dots, X_n \in B_n)$   
 $= \mathbb{P}(X_1 \in B_1) \dots \mathbb{P}(X_n \in B_n)$

for all Borel sets  $B_1, \dots, B_n \in \mathbb{R}$ ,

(iii) the law of  $\underline{X}$  is a product  
 of the laws of  $X_1, \dots, X_n$ , i.e.,

$$\mathbb{P}_{\underline{X}} = \mathbb{P}_{X_1} \times \dots \times \mathbb{P}_{X_n}$$

(iv) the characteristic function of  $\underline{X}$   
 is the product of the characteristic  
 functions of  $X_1, \dots, X_n$ , that is,

$$\mathbb{E}[e^{i(u \cdot \underline{X})}] = \mathbb{E}[e^{iu_1 X_1}] \dots \mathbb{E}[e^{iu_n X_n}]$$

for  $u = (u_1, \dots, u_n) \in \mathbb{R}^h$ .

52) Idea of proof:

(i)  $\leftrightarrow$  (ii) by definition.

(iii)  $\rightarrow$  (ii) clear.

(ii)  $\rightarrow$  (iii): Follows from fact:  
if two Borel prob. measures  $\nu, \mu$   
on  $\mathbb{R}^k$  agree on sets of form  
 $B, x \dots \cup_x B_n \dots, B_i$  Borel, then  
 $\nu = \mu$ .

(iii)  $\rightarrow$  (iv) clear.

(iv)  $\rightarrow$  (iii) follows from fact that  
a measure is uniquely determined by  
its Fourier transform.

Cor. 8.9. If  $X, Y$  are integrable and  
independent, then  
 $E[XY] = E[X] \cdot E[Y]$ .

Proof:  $Z = (X, Y) : \Omega \rightarrow \mathbb{R}^2$ .

$$E[XY] = \int_{\mathbb{R}^2} xy \, dP_Z(x, y)$$

$$\stackrel{\text{Thm. 8.7}}{=} \int_{\mathbb{R}^2} xy \, dP_X \times P_Y(x, y)$$

$$= \left( \int_{\mathbb{R}} x \, dP_X(x) \right) \left( \int_{\mathbb{R}} y \, dP_Y(y) \right)$$

$$= E[X] \cdot E[Y]. \quad \square$$

Lemma 8.10. (Borel-Cantelli II)

$A_n, n \in \mathbb{N}$ , independent events.

53) If  $\sum_{n=1}^{\infty} \mathbb{P}(A_n) = \infty$ , then

$$\mathbb{P}(A_n \text{ i.o.}) = 1.$$

Proof:  $e^{-x} \geq 1-x$  for  $x \in [0, 1]$ .

$$\text{So } \mathbb{P}\left(\bigcup_{n=k}^N A_n\right) = 1 - \mathbb{P}\left(\bigcap_{n=k}^N A_n^c\right)$$

$$\stackrel{\text{indep.}}{=} 1 - \prod_{n=k}^N (1 - \mathbb{P}(A_n))$$

$$\geq 1 - \prod_{n=k}^N e^{-\mathbb{P}(A_n)} = 1 - \exp\left(-\sum_{n=k}^N \mathbb{P}(A_n)\right)$$

$\rightarrow 1$  as  $N \rightarrow \infty$ .

$$\text{So } \mathbb{P}\left(\bigcup_{n \geq k} A_n\right) = 1 \text{ and}$$

$$\mathbb{P}(A_n \text{ i.o.}) = \mathbb{P}\left(\bigcap_k \bigcup_{n \geq k} A_n\right)$$

$$= \lim_{k \rightarrow \infty} \mathbb{P}\left(\bigcup_{n \geq k} A_n\right) = 1. \quad \square$$

Lemma 8.11.  $X, Y: \Omega \rightarrow \mathbb{R}^n$  ind. random var.  $Z = X + Y$ .

Then

and  $\mathbb{P}_Z = \mathbb{P}_X * \mathbb{P}_Y$  (convolution),

$$\phi_Z(u) = \mathbb{E}[e^{i(u \cdot Z)}] = \phi_X(u) \cdot \phi_Y(u) \text{ for } u \in \mathbb{R}^n.$$

Proof:  $\pi: \mathbb{R}^n \times \mathbb{R}^n \rightarrow \mathbb{R}^n$ ,  $\pi(x, y) = x + y$ .

(54) Then  $\mathbb{P}_Z = \pi_* (\mathbb{P}_{(X,Y)})$ .

Since  $X, Y$  are i.i.d.,

$$\mathbb{P}_{(X,Y)} = \mathbb{P}_X \times \mathbb{P}_Y.$$

So if  $A \subseteq \mathbb{R}^n$ , Borel, then

$$\mathbb{P}_Z(A) = \pi_* (\mathbb{P}_{(X,Y)}(A))$$

$$= \int \chi_A \circ \pi \, d\mathbb{P}_{(X,Y)}$$

$$= \int \chi_A(x+y) \, d\mathbb{P}_X(x) \, d\mathbb{P}_Y(y)$$

$$= \int \chi_A \, d(\mathbb{P}_X + \mathbb{P}_Y).$$

Hence  $\mathbb{P}_Z = \mathbb{P}_X + \mathbb{P}_Y$ .

$$\phi_Z(u) = \mathbb{E}[e^{i u \cdot (X+Y)}] = \mathbb{E}[e^{i u \cdot X} \cdot e^{i u \cdot Y}]$$

$$\stackrel{\text{i.i.d.}}{=} \mathbb{E}[e^{i u \cdot X}] \cdot \mathbb{E}[e^{i u \cdot Y}] = \phi_X(u) \cdot \phi_Y(u).$$

## §. 12. Gaussian random variables

$X: \Omega \rightarrow \mathbb{R}$  real-valued random var.

Then  $X$  is Gaussian with mean  $\mu \in \mathbb{R}$  and variance  $\sigma^2 > 0$  if its distribution

$$d\mathbb{P}_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) dx$$

Gaussian or normal distribution!



(55) We write  $X \sim \mathcal{N}(\mu, \sigma^2)$ .

$X$  is standard Gaussian or normal  
if  $X \sim \mathcal{N}(0, 1)$ , i.e.,

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}.$$

If  $X \sim \mathcal{N}(\mu, \sigma^2)$ , then

$$E[X] = \mu, \quad \text{Var}[X] = \sigma^2.$$

$\sigma = \text{Var}[X]^{1/2}$  standard deviation.

characteristic function: if  $X \sim \mathcal{N}(\mu, \sigma^2)$ ,

$$\text{then } \phi_X(u) = \exp\left(-\frac{1}{2} \sigma^2 u^2 + i u \mu\right).$$

If  $X \sim \mathcal{N}(\mu_1, \sigma_1^2)$ ,  $Y \sim \mathcal{N}(\mu_2, \sigma_2^2)$ ,  
and  $X, Y$  are independent, then

$$Z = X + Y \sim \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2).$$

Proof:

$$\begin{aligned} \phi_Z(u) &= \phi_X(u) \cdot \phi_Y(u) \\ &= \exp\left(-\frac{1}{2} (\sigma_1^2 + \sigma_2^2) u^2 + i u (\mu_1 + \mu_2)\right). \quad \square \end{aligned}$$

It is convenient to consider a random  
var.  $X$  s.t.  $X = \mu$  a.s. as a "generalized"  
Gaussian, where  $\sigma^2 = 0$ .

Now,

$$\begin{aligned} \phi_X(u) &= \int_{\mu} \delta(u - \mu) \exp(i u \mu) \\ &= \exp\left(-\frac{1}{2} 0 \cdot u^2 + i u \mu\right). \end{aligned}$$

56) Def. A random variable  $\underline{X} = (X_1, \dots, X_n)$   
 $\Omega \rightarrow \mathbb{R}^n$  is a (generalized, vector-valued) Gaussian iff

$$\phi_{\underline{X}}(u) = \exp\left(-\frac{1}{2} u^t C u + i(u \cdot \mu)\right)$$

$$= E[e^{i(u \cdot \underline{X})}] = \quad \text{for } u \in \mathbb{R}^n$$

where

$\mu \in \mathbb{R}^n$  and  $C$  is a pos. semi-def.  $n \times n$ -matrix

$$\text{Cov}(\underline{X}, \underline{X}) := E[(\underline{X} - E[\underline{X}])(\underline{X} - E[\underline{X}])^t]$$

covariance of  $\underline{X}, \underline{X}$

Then  $C$  is the covariance matrix of  $\underline{X}$ , i.e.  $C = (c_{ij})$ , where

$$c_{ij} = \text{Cov}(X_i, X_j)$$

$\underline{X}$  Gaussian iff  $\underline{X} = B \underline{Y}$ ,

where  $B$   $n \times n$ -matrix,  $\underline{Y} = (Y_1, \dots, Y_n)$

s.t.  $Y_1, \dots, Y_n$  real-valued ind.

generalized Gaussians iff

$$\underline{X} = D \underline{Z} + a, \text{ where } a \in \mathbb{R}^n, D: h \times k \text{-matrix,}$$

$$\underline{Z} = (Z_1, \dots, Z_k), Z_1, \dots, Z_k \text{ ind. Gaussians.}$$

$$\underline{Y} = A \underline{X}, \quad A: n \times k \text{ matrix.}$$

$$\underline{X}: \Omega \rightarrow \mathbb{R}^k, \quad \underline{Y}: \Omega \rightarrow \mathbb{R}^n$$

If  $\underline{X}$  Gaussian, then  $\underline{Y}$  is Gaussian.

(57) Proof:  $\phi_{\underline{X}}(\underline{v}) = \mathbb{E}[e^{i(\underline{v} \cdot \underline{X})}]$

$$= \mathbb{E}[e^{i(\underline{v} \cdot A\underline{X})}] = \mathbb{E}[e^{i(A^t \underline{v} \cdot \underline{X})}]$$

$$= \phi_{\underline{X}}(\underbrace{A^t \underline{v}}_{\underline{u}}) = \exp\left(-\frac{1}{2} (A^t \underline{v})^t C A^t \underline{v} + i (A^t \underline{v} \cdot \underline{\mu})\right)$$

$$= \exp\left(-\frac{1}{2} \underline{v}^t (A C A^t) \underline{v} + i (\underline{v} \cdot A \underline{\mu})\right)$$

S.  $\underline{\mu}' = A \underline{\mu}$ ,  $C' = A C A^t$ .  $\square$

If  $\underline{X}: \Omega \rightarrow \mathbb{R}^n$  has a multi-modal distribution given by

$$dP_{\underline{X}}(\underline{x}) = \frac{|A|^{1/2}}{(2\pi)^{n/2}} \exp\left(-\frac{1}{2} (\underline{x} - \underline{\mu})^t A (\underline{x} - \underline{\mu})\right),$$

where  $\underline{\mu} \in \mathbb{R}^n$ , and  $A$  pos. def.  $n \times n$ -matrix,

then  $\underline{X}$  is Gaussian and

$$\phi_{\underline{X}}(\underline{u}) = \exp\left(-\frac{1}{2} \underline{u}^t C \underline{u} + i (\underline{u} \cdot \underline{\mu})\right),$$

where  $C = A^{-1}$ .

Ex. 13. Modes of convergence of random variables

$\underline{X}_n$ ,  $n \in \mathbb{N}$  ( $\omega$ ), real (or vector valued) random variables

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$X_n \rightarrow X_\infty$  a.s. (almost surely)

iff  $\mathbb{P}(X_n \rightarrow X_\infty) =$

$$\mathbb{P}(\{\omega \in \Omega: X_n(\omega) \rightarrow X(\omega)\}) = 1$$

iff  $X_n(\omega) \rightarrow X_\infty(\omega)$  for a.e.  $\omega \in \Omega$ .

$X_n \rightarrow X_\infty$  in probability

$$\text{iff } \lim_{n \rightarrow \infty} \mathbb{P}(|X_n - X_\infty| \geq \epsilon) = 0$$

for all  $\epsilon > 0$ .

(equiv. to "convergence in measure")

$X_n \rightarrow X_\infty$  in  $L^p$ ,  $p \geq 1$ , iff

$$\mathbb{E}[|X_n - X_\infty|^p] \rightarrow 0$$

equiv.  $\int_{\Omega} |X_n(\omega) - X_\infty(\omega)|^p d\mathbb{P}(\omega) \rightarrow 0$ .

$X_n \rightarrow X_\infty$  a.s.

$\exists$  subseq.

$X_n \rightarrow X_\infty$   
in prob.

$X_n \rightarrow X_\infty$  in  $L^p$

Proof (easy) e.g. 1: Fix  $\epsilon > 0$

$$E_n = \{\omega \in \Omega: |X_n(\omega) - X_\infty(\omega)| \geq \epsilon\}$$

(59)  $\underline{X}_n \rightarrow \underline{X}_\infty$  a.s. implies

$$0 = \mathbb{P}(E_n \text{ i.e.}) = \mathbb{P}\left(\bigcap_{n \geq k} \bigcup_{n \geq k} E_n\right)$$

$$\leq \lim_{n \rightarrow \infty} \mathbb{P}\left(\bigcup_{n \geq k} E_n\right) \leq \limsup_{n \rightarrow \infty} \mathbb{P}(E_n),$$

Lem. 8.14.  $\underline{X}_n$   $\mathbb{R}^d$ -valued Gaussian random variables,  $n \in \mathbb{N}$ .

$\underline{X}_n \rightarrow \underline{X}_\infty$  in probability.

Then  $\underline{X}_\infty$  is  $\mathbb{R}^d$ -valued Gaussian.

Proof (outline): 1. If  $\underline{X}_n \rightarrow \underline{X}_\infty$  in prob., then  $\phi_{\underline{X}_n}(u) \rightarrow \phi_{\underline{X}_\infty}(u)$  locally uniformly on  $\mathbb{R}^d$ .

$$\begin{aligned} |e^{i(u \cdot \underline{X}_n)} - e^{i(u \cdot \underline{X}_\infty)}| &\leq |(u \cdot \underline{X}_n) - (u \cdot \underline{X}_\infty)| \\ &\leq |u| \cdot |\underline{X}_n - \underline{X}_\infty| \end{aligned}$$

$$\text{So } \left| \phi_{\underline{X}_n}(u) - \phi_{\underline{X}_\infty}(u) \right|$$

$$\leq \mathbb{E} \left[ \left| e^{i(u \cdot \underline{X}_n)} - e^{i(u \cdot \underline{X}_\infty)} \right| \right]$$

$$\leq |u| \cdot \sqrt{\epsilon} + 2 \cdot \mathbb{P}(|\underline{X}_n - \underline{X}_\infty| \geq \sqrt{\epsilon}) \leq \epsilon$$

(\*) follows. for  $\epsilon$  large.

2.  $\underline{X}_n$  Gaussian; so

$$\phi_{\underline{X}_n}(u) = \exp\left(-\frac{1}{2} u^T C_n u + i(u \cdot \mu_n)\right)$$

(60) If  $\mu \in \mathbb{R}^d$   
 $\phi_{\underline{X}_n}(u) \longrightarrow \phi_{\underline{X}}(u)$  loc. unid.,  
 then  $\phi_{\underline{X}}$  has the same form, i.e.,  
 $\phi_{\underline{X}}(u) = \exp\left(-\frac{1}{2} u^T C u + i(u \cdot \mu)\right)$ ,  
 where  $C \geq 0$ ,  $\mu \in \mathbb{R}^d$ .  $\square$

Lev. P. 15.  $X_1, \dots, X_n$  real-valued random var. with joint Gaussian distribution (i.e.,  $\underline{X} = (X_1, \dots, X_n)$  is  $\mathbb{R}^n$ -valued Gaussian random var.)

Then  $X_1, \dots, X_n$  are independent iff they are pairwise uncorrelated, i.e.

$$\text{Cov}(X_i, X_j) = 0 \text{ for } i, j = 1, \dots, n, i \neq j.$$

Proof:  $\Rightarrow$  Clear:

$$\text{Cov}(X_i, X_j) = \mathbb{E}[(X_i - \mathbb{E}[X_i])(X_j - \mathbb{E}[X_j])] \\
\stackrel{X_i, X_j \text{ ind.}}{=} \underbrace{\mathbb{E}[X_i - \mathbb{E}[X_i]]}_0 \cdot \underbrace{\mathbb{E}[X_j - \mathbb{E}[X_j]]}_0 = 0$$

$\Leftarrow$  Since  $\underline{X}$  is Gaussian,

$$\phi_{\underline{X}}(u) = \exp\left(-\frac{1}{2} u^T C u + i(u \cdot \mu)\right), \quad u \in \mathbb{R}^n,$$

where  $C = (C_{ij})$  is the covariance matrix;

$$\text{so } C_{ij} = \text{Cov}(X_i, X_j), \quad i, j = 1, \dots, n.$$

(61) By assumption  $c_{ij} = 0$  for  $i \neq j$ ,  
and so  $C$  is a diagonal matrix.

Hence

$$\phi_{\underline{X}}(u) = \phi_{\underline{X}_1}(u_1) \cdots \phi_{\underline{X}_n}(u_n),$$

for  $u = (u_1, \dots, u_n) \in \mathbb{R}^n$ .

This shows that  $\underline{X}_1, \dots, \underline{X}_n$  are independent.

by Thm. 8.4.  $\square$

### §. 16. Stochastic processes

A stochastic process in  $\mathbb{R}^n$  is a collection  
 $\{\underline{X}_t\}_{t \in T}$  of random variables defined

on a probability space  $(\Omega, \mathcal{A}, \mathbb{P})$ ,

where  $T \subseteq \mathbb{R}$  "parameter set of "times",

$T = \mathbb{N}_0, \mathbb{N}$  "discrete time stochastic process";

Sequence of random variables:

$$\underline{X}_0, \underline{X}_1, \dots$$

$$T = [0, \infty), [a, b], \text{ etc.}$$

"continuous time stochastic process".

$t \in T$  fixed:  $\omega \mapsto \underline{X}_t(\omega)$  random var.  
on  $\Omega$ .

$\omega$  fixed:  $t \in T \mapsto \underline{X}(t, \omega) := \underline{X}_t(\omega)$   
sample path of stochastic process.

⑥2 Def. 1.71 (Brownian motion)

A real-valued stochastic process  $\{B_t\}_{t \in [0, \infty)}$  is called a (version of) Brownian motion if the following conditions are true:

(i) the process is a Gaussian process, i.e., for all  $n \in \mathbb{N}$ ,  $0 \leq t_1 < \dots < t_n$  the random variables  $B_{t_1}, \dots, B_{t_n}$  have a joint Gaussian distribution.

(ii)  $B_t$  for  $t \in [0, \infty)$  is centered,  
i.e.  $E[B_t] = 0$

(iii)  $\text{Cov}(B_s, B_t) = E[B_s \cdot B_t]$   
 $= s \wedge t$ ,  $s, t \in [0, \infty)$ ,

(iv) sample paths  $t \mapsto B_t$  are continuous a.s.,  
(i.e.,  $t \mapsto B_t(\omega)$  is cont. for a.e.  $\omega$ ).

Rev. 1.18.  $\{B_t\}_{t \in [0, \infty)}$  a Brownian motion.

$$1) E[B_t] = 0, \text{Var}(B_t) = \text{Cov}(B_t, B_t) = t \text{ for } t \geq 0.$$

$$B_t \sim \mathcal{N}(0, t), \quad t \geq 0$$

$$B_0 = 0 \text{ a.s.}$$

Brownian motion starts at 0 for time 0 a.s.

(63) 2) Brownian motion has "independent increments":

If  $t_1 < t_2 < \dots < t_n$ , then

(\*)  $\bar{X}_{t_1}, \bar{X}_{t_2} - \bar{X}_{t_1}, \bar{X}_{t_3} - \bar{X}_{t_2}, \dots, \bar{X}_{t_n} - \bar{X}_{t_{n-1}}$   
are ind. Gaussian rand. varb.

$$\bar{X}_{t_k} - \bar{X}_{t_{k-1}} \sim \mathcal{N}(0, t_k - t_{k-1})$$

Indeed: r.v.'s in (\*) joint Gaussian,  
centered, and

$k < l$ .

$$t_{k-1} \leq t_k \leq t_{l-1} < t_l$$

$$\text{Cov}(\bar{X}_{t_k} - \bar{X}_{t_{k-1}}, \bar{X}_{t_l} - \bar{X}_{t_{l-1}})$$

$$= \mathbb{E}[(\bar{X}_{t_k} - \bar{X}_{t_{k-1}})(\bar{X}_{t_l} - \bar{X}_{t_{l-1}})]$$

$$= t_k \wedge t_l - t_{k-1} \wedge t_l - t_k \wedge t_{l-1} + t_{k-1} \wedge t_{l-1}$$

$$= t_k - t_{k-1} - t_k + t_{k-1} = 0.$$

So by Lev. 8.15, the r.v.'s in (\*) are ind.

## §.19. Hilbert space bases

real

$H$  separable Hilbert space,

A sequence  $\{x_n\}_{n \in \mathbb{N}}$  is called a complete orthonormal system or a

Hilbert space basis

if i) the vectors  $\{x_n\}_{n \in \mathbb{N}}$  are orthonormal, i.e.,  $(x_i, x_j) = \delta_{ij}$   $i, j \in \mathbb{N}$ ,

(64) ii) if  $x \in H$  and  $(x, x_n) = 0$  for all  $n \in \mathbb{N}_0$ , then  $x = 0$ .

In this case,

$$x = \sum_{n=1}^{\infty} (x, x_n) x_n.$$

$$\|x\|^2 = \sum_{n=1}^{\infty} |(x, x_n)|^2$$

Parseval's Identities.

$$(x, y) = \sum_{n=1}^{\infty} (x, x_n) \cdot (y, x_n)$$

Equivalent to (ii) is

(ii'): The set  $S$  of all (finite) linear combinations of the vectors  $x_1, x_2, x_3, \dots$  is dense in  $H$ .

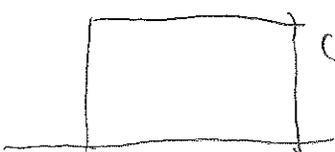
Ex.  $H = L^2[0, 1]$   $(f, g) = \int_0^1 f(x)g(x)dx$

Hilbert space bases: trigonometric functions

$$1, \frac{1}{\sqrt{2}} \cos(2\pi n x), \frac{1}{\sqrt{2}} \sin(2\pi n x), n \in \mathbb{N}$$

2. Haar basis  $n \in \mathbb{N}_0$

$$\varphi_{n,k}(x) := \begin{cases} 1 & \text{for } x \in \left[\frac{k}{2^n}, \frac{k+1/2}{2^n}\right) \\ -1 & \text{for } x \in \left[\frac{k+1/2}{2^n}, \frac{k+1}{2^n}\right) \\ 0 & \text{else } k=0, \dots, 2^n-1 \end{cases}$$

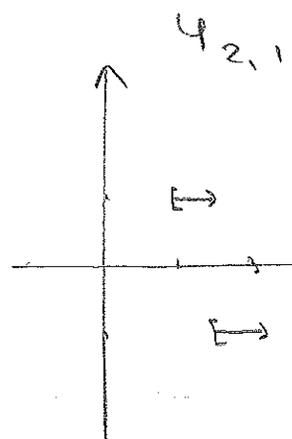
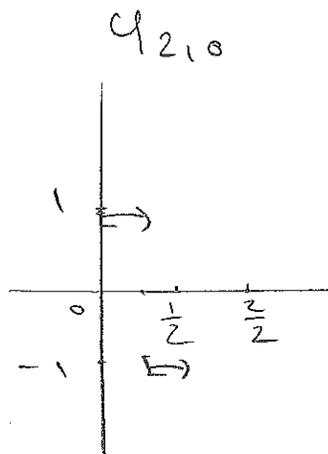
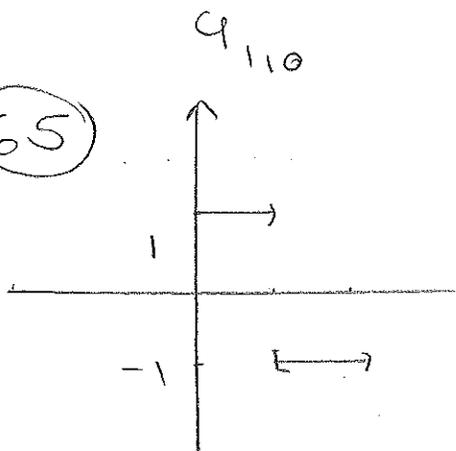


$\varphi_{-1,0} \equiv 1$

$\varphi_{1,0}$  else  $k=0, \dots, 2^n-1$

$I$  set of indices

(65)



Obviously,  $\varphi_{n,k} \in L^2[0, 1]$  pairwise orthogonal

$$\|\varphi_{n,k}\|^2 = \int_0^1 \varphi_{n,k}^2(x) dx = \frac{1}{2^n} \quad n \in \mathbb{N}$$

$$\psi_{n,k} = 2^{n/2} \varphi_{n,k} \quad \psi_{1,0} \equiv 1$$

$\{\psi_{n,k}\}_{(n,k) \in \mathbb{I}}$  orthonormal system.

Linear comb. are dense in  $L^2[0, 1]$ :  
(because step functions on dyadic intervals are);

So  $\{\psi_{n,k}\}_{(n,k) \in \mathbb{I}}$  is a Hilbert space basis of  $L^2[0, 1]$ .

If  $\{x_n\}_{n \in \mathbb{N}_\infty}$  is an orthonormal system, then  $\sum_{n=1}^{\infty} a_n x_n$  converges

$$\text{iff } \sum_{n=1}^{\infty} a_n^2 < \infty$$

Follows from Cauchy criterion:

$$S_n = \sum_{k=1}^n a_k x_k \text{ partial sum}$$

$$(66) \quad \|S_n - S_m\|^2 = \sum_{k=m+1}^n a_k^2, \quad n \geq m.$$

## § 20. Construction of Brownian motion

1. on  $T = [0, 1]$

$Z_n, n \in \mathbb{N}$ , i.i.d. r.v.'s  
independent, identically distributed

t.v.'s on same p.v.b. space  
 $(\Omega, \mathcal{A}, \mathbb{P}) \quad Z_n \sim \mathcal{N}(0, 1)$

$\tilde{\Omega} = (\mathbb{R}, \mathcal{B}, \mu)$

$\mathcal{B}$  Borel  $\sigma$ -alg.  $d\mu(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$

$\tilde{\Omega} \cong \mathbb{R}^{\mathbb{N}}$ ,  $Z_n =$  projection onto  $n$ -th  
coordinate

$Z_n, n \in \mathbb{N}$ , orthonormal system  
in  $L^2(\tilde{\Omega})$ :

$$\int_{\tilde{\Omega}} Z_n(\omega) \cdot Z_k(\omega) d\mathbb{P}(\omega) = \text{cov}(Z_n, Z_k) = \delta_{nk}, \quad n, k \in \mathbb{N}.$$

$\varphi_n$  are  $\mathcal{N}$  orthonormal Hilbert space basis  
at  $L^2[0, 1]$ .

$$f_n(t) = \int_0^t \varphi_n(t) dt = (\varphi_n, \chi_{[0, t]})$$

inner product in  $L^2[0, 1]$ .

$$\boxed{\mathcal{B}_t = \sum_{n=1}^{\infty} f_n(t) Z_n \quad f_n, t \in [0, 1]}$$

(67) i) For each  $t \in [0,1]$  the sum converges in  $L^2(\Omega)$    
 equi- $\cup$

$$\sum_{n=1}^{\infty} f_n(t)^2 = \sum_{n=1}^{\infty} (\chi_{[0,t]})^2$$

$$= \|\chi_{[0,t]}\|_{L^2[0,1]}^2 = t < \infty.$$

Parseval

ii) each  $B_t$  is a Gaussian; actually, for  $t_1 < t_2 < \dots < t_m$ ,

$B_{t_1}, B_{t_2}, \dots, B_{t_m}$  have a joint Gaussian distribution.

$$B_t^n := \sum_{k=1}^n f_k(t) Z_k \text{ is Gaussian (linear comb. of Gaussians).}$$

$B_t^n \rightarrow B_t$  as  $n \rightarrow \infty$  in  $L^2(\Omega)$ ,  
 so  $B_t$  is Gaussian by Lec. 8.14.

Similarly,

$(B_{t_1}^n, \dots, B_{t_m}^n)$  have a joint Gaussian distribution, and

$$(B_{t_1}^n, \dots, B_{t_m}^n) \rightarrow (B_{t_1}, \dots, B_{t_m})$$

as  $n \rightarrow \infty$  in  $L^2(\Omega, \mathbb{R}^m)$ ;

so  $(B_{t_1}, \dots, B_{t_m})$  have a joint Gaussian distribution.

(68) (iii)  $B_t$  is centered:

$$\begin{aligned} \mathbb{E}[B_t] &= \int_{\Omega} B_t(\omega) d\mathbb{P}(\omega) \\ &= \lim_{n \rightarrow \infty} \int_{\Omega} B_t^n(\omega) d\mathbb{P}(\omega) \\ &= 0, \text{ because } Z_n, n \in \mathbb{N}, \text{ is centered.} \\ &= \lim_{n \rightarrow \infty} \sum_{k=1}^n f_k(t) \mathbb{E}[Z_k] = 0. \end{aligned}$$

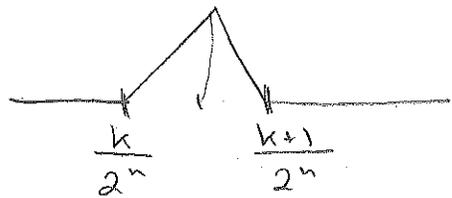
$$\begin{aligned} \text{(iv) Cov}(B_s, B_t) &= \int_{\Omega} B_s(\omega) B_t(\omega) d\mathbb{P}(\omega) \\ &= \lim_{n \rightarrow \infty} \int_{\Omega} B_s^n(\omega) B_t^n(\omega) d\mathbb{P}(\omega) \\ &= \lim_{n \rightarrow \infty} \sum_{k, l=1}^n f_k(s) f_l(t) \underbrace{\text{Cov}(Z_k, Z_l)}_{\delta_{kl}} \\ &= \sum_{k=1}^{\infty} f_k(s) f_k(t) \\ &= \sum_{k=1}^{\infty} (\varphi_{k,1} \chi_{[0,s]}) (\varphi_{k,1} \chi_{[0,t]}) \end{aligned}$$

$$\stackrel{\text{Parseval}}{=} (\chi_{[0,s]} | \chi_{[0,t]}) = s \wedge t.$$

To check continuity at  $t \rightarrow B_t(\omega)$   
 For a.e.  $\omega \in \Omega$ , we choose the  
 Haar basis for the Hilbert space  
 basis of  $L^2[0,1]$ .

(69)

$$f_{n,k}(t) = \int_0^t \varphi_{n,k}(x) dx \quad \text{Lip}(f_{n,k}) = 2^{n/2}$$



$$\|f_{n,k}\|_\infty \leq \frac{1}{2} \cdot \frac{1}{2^n} \cdot 2^{n/2} = \frac{1}{2^{n/2+1}}$$

$$\approx 2^{-n/2}$$

Claim  $\{Z_{n,k}\}_{(n,k) \in \mathbb{I}}$  i.i.d. standard normal r.v.'s.

Then for a.e.  $\omega \in \Omega$ ,

the series  $g_{-1}(t)$

$$B_t(\omega) = \underbrace{Z_{-1,0}(\omega)}_{g_{-1}(t)} f_{-1}(t) + \sum_{n=0}^{\infty} \underbrace{\sum_{k=0}^{2^n-1} Z_{n,k}(\omega) f_{n,k}(t)}_{g_n(t)}$$

converges uniformly in  $t$  (and hence represents a cont. function in  $t$ ).

Proof:  $\frac{2}{\sqrt{2\pi}} \int_a^\infty e^{-x^2/2} dx \leq e^{-a^2/2}$  for  $a > 0$ .

$\therefore \mathbb{P}(|Z| > a) \leq e^{-a^2/2}$  for  $a \geq 0$ , if  $Z \sim \mathcal{N}(0,1)$ .

$$A_{n,k} = \{|Z_{n,k}| > 2\sqrt{\log(2^{n/2}/n)}\}$$

$$\mathbb{P}(A_{n,k}) \leq \exp(-2 \log(2^{n/2}/n)) = \frac{1}{2^n \cdot n^2}$$

$$\textcircled{70} \quad \sum_{n=1}^{\infty} \sum_{k=0}^{2^n-1} \mathbb{P}(A_{n,k}) \leq \sum_{n=1}^{\infty} \frac{2^n}{2^n - n^2}$$

$$\leq \sum_{n=1}^{\infty} \frac{1}{n^2} < \infty$$

So, by Borel-Cantelli I, we have

$$\mathbb{P}(A_{n,k} \text{ i.o.}) = 0 \quad \text{i.e., for a.e.}$$

$\omega \in \Omega$  we have

$$|Z_{n,k}(\omega)| \leq 2 \cdot \sqrt{1_{A_{n,k}}(\omega)} \leq \sqrt{n} \cdot 1_{A_{n,k}}(\omega)$$

(\*) For all suff. large  $n$   
(dep. on  $\omega$ ), so for  $n \geq N(\omega)$ .

For such  $\omega$ :

$$\sum_{n=N(\omega)}^{\infty} \left| \sum_{k=0}^{2^n-1} \underbrace{Z_{n,k}(\omega) \cdot \mathbb{1}_{A_{n,k}}(\omega)}_{= g_n(\omega)} \right|$$

$$\leq \sum_{n=N(\omega)}^{\infty} \frac{\sqrt{n}}{2^{n/2}} < \infty \quad ; \text{ so series}$$

(\*) represents a cont. function in  $t$  by Weierstrass - M-test.

Actually,

$g_n$  is  $L_n$ -Lipschitz  $n^{1/2}$   
with  $L_n \leq \sqrt{n} \cdot 2$

for all  $n \geq N(\omega)$ ; so

by adjusting constants wlog  
for all  $n \geq 1$

(7) Moreover,

$$\|g_n\|_\infty \leq \sqrt{n} \cdot 2^{-n/2} \quad \text{for all } n \geq N(\omega),$$

w.l.g. for all  $n \geq 1$ .

Suppose  $\omega$  is "good"  $\omega$ ; so that it satisfies  $\dagger \dagger$ .

Let  $s, t \in [0, 1]$ . Pick suitable  $N = N(s, t) \in \mathbb{N}$ .

Then

$$|B_s(\omega) - B_t(\omega)|$$

$$\leq \sum_{n=-1}^{\infty} |g_n(s) - g_n(t)|$$

$$\leq \sum_{n=-1}^{N(s,t)} L_n |s-t| + \sum_{n=N+1}^{\infty} 2 \cdot \|g_n\|_\infty$$

$$\stackrel{2.3}{\leq} 1 + \sum_{n=1}^N \sqrt{n} 2^{n/2} |s-t| + \sum_{n=N}^{\infty} \sqrt{n} 2^{-n/2}$$

$$\stackrel{2.3}{\leq} \sqrt{N} \cdot 2^{N/2} |s-t| + \sqrt{N} \cdot 2^{-N/2}$$

Pick  $N = N(s, t)$  s.t.

$$2^{N/2} |s-t| \approx 2^{-N/2} \quad \text{equiv. } |s-t| \approx 2^{-N}$$

$$\text{equiv. } N = \log_2 \frac{1}{|s-t|} \approx \log \frac{1}{|s-t|}.$$

Then

$$|B_s(\omega) - B_t(\omega)| \leq |s-t|^{1/2} \sqrt{\log \frac{1}{|s-t|}}$$

72 Conclusion For a.e.  $\omega$  there ex.  
 $M(\omega) \geq 0$  s.t.

$$|B_s(\omega) - B_t(\omega)| \leq M(\omega) |s-t|^{1/2} \sqrt{\log \frac{1}{|s-t|}}$$

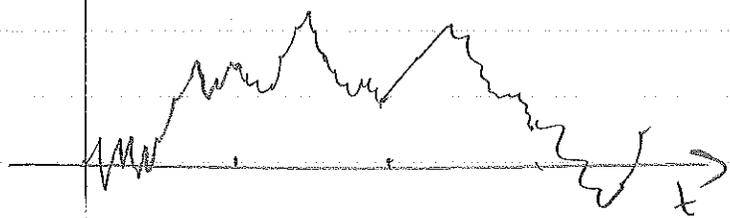
Almost surely, the sample path  $t \mapsto B_t(\omega)$   
 has modulus of continuity

$$\omega(\delta) = C \delta^{1/2} \sqrt{\log \frac{1}{\delta}};$$

So for every  $\epsilon > 0$ ,  $t \mapsto B_t(\omega)$   
 is Hölder  $(1/2 - \epsilon)$  almost surely.

## 2. Brownian motion on $[0, \infty)$

Idea: Let a Brownian motion run  
 until time 1, start a "new" Brownian  
 at endpoint, let it run until time 2,  
 etc.  $B_t(\omega)$   $\omega$  fixed



$B_t$ ,  $n \in \mathbb{N}_0$ , independent copies  
 of Brownian motion on  $[0, 1]$ .

Define

$$B_t(\omega) = \sum_{k=0}^{\lfloor t \rfloor - 1} B_1^k(\omega) + B_{t - \lfloor t \rfloor}^{\lfloor t \rfloor}(\omega)$$

$$B_{1.5}(\omega) = B_1^0(\omega) + B_{0.5}^1(\omega).$$

73) Then  $\{B_t\}_{t \in [0, \infty)}$  is a Gaussian process.

$B_t$  centered,  $s \leq t$

$$\text{Cov}(B_s, B_t) =$$

$$E \left[ \left( \sum_{k=0}^{\lfloor s \rfloor - 1} B_k + B_{s - \lfloor s \rfloor} \right) \left( \sum_{l=0}^{\lfloor t \rfloor - 1} B_l + B_{t - \lfloor t \rfloor} \right) \right]$$

$$= \sum_{k=0}^{\lfloor s \rfloor - 1} 1 + (s - \lfloor s \rfloor) = s = s \wedge t, \quad \lfloor s \rfloor \leq \lfloor t \rfloor$$

For each  $n \in \mathbb{N}_0$ :  $t \mapsto B_t^n(\omega) \in [0, 1]$   
 cont. a.s. so for e.e.  $\omega$ :  
 $t \mapsto B_t^n(\omega)$  cont. for all  $n \in \mathbb{N}_0$ .

Hence  $t \mapsto B_t(\omega)$  cont. a.s.

### §. 21. $\pi$ -systems

$\underline{X}$  set,  $\mathcal{G}$  family of subsets of  $\underline{X}$ .

$\mathcal{G}$  is a  $\pi$ -system if

$A \cap B \in \mathcal{G}$  whenever  $A, B \in \mathcal{G}$ .

(i.e., a  $\pi$ -system is "stable" under finite intersections.)

Facts 1)  $\mathcal{G}$   $\pi$ -system,  $\mathcal{A} = \sigma(\mathcal{G})$   $\sigma$ -algebra generated by  $\mathcal{G}$ ,  $\mu, \nu$  prob. measures on  $\mathcal{A}$ .

$\mu(A) = \nu(A)$  for all  $A \in \mathcal{G}$ ,

then  $\mu = \nu$  (i.e.,  $\mu(A) = \nu(A)$  for all  $A \in \mathcal{A}$ .)

74 Exercise 1  $(\Omega, \mathcal{A}, \mathbb{P})$  prob. space

2)  $\mathcal{G}, \mathcal{F}$   $\pi$ -systems,  $\mathcal{B} = \sigma(\mathcal{G})$ ,  $\mathcal{E} = \sigma(\mathcal{F})$ ,  
in  $\mathcal{A}$ .

$$\text{If } \mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

whenever  $A \in \mathcal{G}$ ,  $B \in \mathcal{F}$ , then  $\mathcal{B}$  and  $\mathcal{E}$   
are independent, i.e.

$$\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

for all  $A \in \mathcal{B}$ ,  $B \in \mathcal{E}$ .

Exercise!

8.22. The space  $X = C([0, \infty))$ .

$X := C([0, \infty)) = \{f: [0, \infty) \rightarrow \mathbb{R} \text{ cont.}\}$   
equipped with topology of loc. unif. convergence

$$f_n \rightarrow f \text{ iff } f_n \rightarrow f \text{ loc. unif. on } \mathbb{R}$$

This is a metrizable topology

$$d_n(f, g) = \sup_{x \in [0, n]} |f(x) - g(x)|$$

$$d(f, g) = \sum_{n=1}^{\infty} \frac{1}{2^n} \underbrace{\frac{d_n(f, g)}{1 + d_n(f, g)}}_{\leq 1}$$

$d$  metric on  $X$

$$d(f_n, f) \rightarrow 0 \text{ iff } f_n \rightarrow f \text{ loc. unif. on } [0, \infty).$$

$(X, d)$  separable metric space.

(75)  $\mathcal{B} = \mathcal{B}_{\mathbb{X}}$  Borel  $\sigma$ -algebra on  $\mathbb{X}$   
 (= smallest  $\sigma$ -algebra containing all  
 open sets in  $\mathbb{X}$ )

Want to find  $\pi$ -system  $\mathcal{G}$  s.t.

$$\mathcal{B} = \sigma(\mathcal{G})!$$

$$t \in [0, \infty)$$

$$\pi_t : \mathbb{X} \longrightarrow \mathbb{R} = \text{evaluation at time } t$$

$$f \longmapsto f(t)$$

$$\mathcal{G} = \left\{ \pi_{t_1}^{-1}(B_1) \cap \dots \cap \pi_{t_k}^{-1}(B_k) : k \in \mathbb{N}, \right. \\ \left. t_1 < \dots < t_k \text{ in } [0, \infty), \right. \\ \left. B_1, \dots, B_k \in \mathcal{B}_{\mathbb{R}} \right\}.$$

Obviously,  $\mathcal{G}$  is a  $\pi$ -system!

Claim  $\sigma(\mathcal{G}) = \mathcal{B}_{\mathbb{X}}$

Proof (Outline): 1. For  $t \in [0, \infty)$

$\pi_t : \mathbb{X} \rightarrow \mathbb{R}$  is cont. i.e.  $\pi_t^{-1}(B) \in \mathcal{B}_{\mathbb{X}}$   
 for each  $B \in \mathcal{B}_{\mathbb{R}}$ , and  $\mathcal{G} \subseteq \mathcal{B}_{\mathbb{X}}$ .  
 Hence  $\sigma(\mathcal{G}) \subseteq \mathcal{B}_{\mathbb{X}}$ .

2.  $\mathcal{B}_{\mathbb{X}} \subseteq \sigma(\mathcal{G})$ .

Let  $f_0 \in \mathbb{X}$  be arb.

$$f \longmapsto |f(t) - f_0(t)| \text{ is } \sigma(\mathcal{G})\text{-measurable}$$

$$f \longmapsto d_n(f, f_0) = \sup_{t \in [0, n] \cap \mathbb{Q}} |f(t) - f_0(t)| \text{ is } \sigma(\mathcal{G})\text{-measurable}$$

$$(76) \quad f \mapsto d(f, f_0) = \sum_{n=1}^{\infty} \frac{1}{2^n} \frac{d_n(f, f_0)}{1 + d_n(f, f_0)}$$

is  $\mathcal{B}(\mathcal{F})$ -measurable.

Open balls  $B_d(f_0, \epsilon) = \{f : d(f, f_0) < \epsilon\}$   
are  $\mathcal{B}(\mathcal{F})$ -measurable.

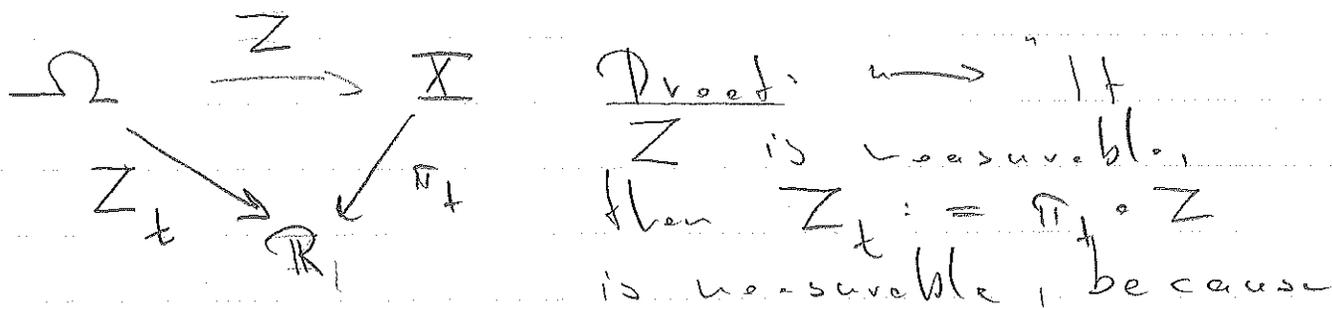
Since every open set in  $\bar{X}$  is a countable union of open balls, every open set is in  $\mathcal{B}(\mathcal{F})$ .

Hence  $\mathcal{B}_{\bar{X}} \subseteq \mathcal{B}(\mathcal{F})$ .

$(\Omega, \mathcal{A}, \mathbb{P})$  prob. space

$Z: \Omega \rightarrow \bar{X}$  is measurable (w.r.t.  $\mathcal{A}$  and  $\mathcal{B}_{\bar{X}}$ )  
iff Claim

$Z_t := \pi_t \circ Z$  is measurable  
for each  $t \in [0, \infty)$ .



$\pi_t \circ Z_t$  is cont.  
 $\mathcal{E} = \{A \subseteq \bar{X} : Z^{-1}(A) \in \mathcal{A}\}$   
 $\mathcal{E}$   $\sigma$ -algebra

$B \in \mathcal{B}_{\mathbb{R}}$ ,  $t \in [0, \infty)$

$$Z^{-1}(\pi_t^{-1}(B)) = (\pi_t \circ Z)^{-1}(B) = Z_t^{-1}(B) \in \mathcal{A}_t$$

so  $\pi_t^{-1}(B) \in \mathcal{E}$ . Hence  
 $\mathcal{G} \subseteq \mathcal{E}$  and  $\mathcal{B}(\mathcal{F}) = \mathcal{B}_{\bar{X}} \subseteq \mathcal{E}$ .  $\square$

(77) Thm. §.23. (Canonical Brownian motion)  
 $X = C([0, \infty))$ ,  $\mathcal{B} = \mathcal{B}_X$  Borel  $\sigma$ -algebra on  $X$ . There exists a unique probability measure  $W$  on  $(X, \mathcal{B})$ , called Wiener measure with the following property:

(i) If we define  $B_t = \pi_t^{-1}(0)$  then  $\{B_t\}_{t \in [0, \infty)}$  is a Brownian motion (on  $\mathbb{R}$ ); were explicitly  $\mathbb{R}$

(i) For  $t_1 < \dots < t_k$  with the random var.  $B_{t_1}, \dots, B_{t_k}$  have a joint Gaussian distribution;

equiv. = Let  $F \subseteq [0, \infty)$  be finite  $|F|$   
 $\pi_F: X \rightarrow \mathbb{R}^F = \{ \varphi: F \rightarrow \mathbb{R} \} \cong \mathbb{R}^{|F|}$   
 $\mu_F = (\pi_F)_* (W)$  is a Gaussian measure on  $\mathbb{R}^F$   
 $\mu_t = (\pi_t)_* (W)$

(ii)  $B_t$  is centered equiv.

$$\int_{\mathbb{R}} x d\mu_t(x) = 0 \quad \text{for each } t \in \mathbb{R}$$

(iii) Cov  $(B_s, B_t) = s \wedge t$  equiv.

$$\int_{\mathbb{R}^2} xy d\mu_{s \wedge t}(x, y) = s \wedge t.$$

Proof: 1. Uniqueness: Suppose  $W, \tilde{W}$  are two measures with the properties (i) - (iii).

Then  $(\pi_F)_* (W) = \mu_F = \tilde{\mu}_F = (\pi_F)_* (\tilde{W})$ ,  
 for each  $F \subseteq [0, \infty)$

because the Fourier transforms of  $\mu_F, \tilde{\mu}_F$  finite

(78) and hence  $\mu_F, \tilde{\mu}_F$  themselves are uniquely determined by (i) - (iii).

This implies that for  $t_1 \leq \dots \leq t_n$ ,  $B_1, \dots, B_n \in \mathcal{B}_{\mathbb{R}}$  we have  $F = \{t_1, \dots, t_n\}$

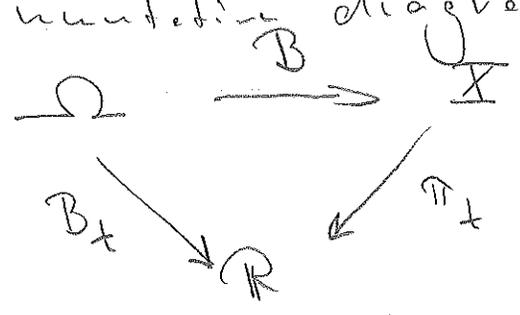
$$\begin{aligned} & W(\pi_{t_1}^{-1}(B_1) \cap \dots \cap \pi_{t_n}^{-1}(B_n)) \\ &= W(\pi_F^{-1}(B_1 \times \dots \times B_n)) = \mu_F(B_1 \times \dots \times B_n) \\ &= \tilde{W}(\tilde{\pi}_{t_1}^{-1}(B_1) \cap \dots \cap \tilde{\pi}_{t_n}^{-1}(B_n)), \end{aligned}$$

i.e.,  $W(S) = \tilde{W}(S)$  for all  $S \in \mathcal{G}$ .  
 Since  $\mathcal{G}(S) = \mathcal{B}_{\mathbb{X}}$ , we have  $W = \tilde{W}$ .

2. Existence: There ex. Brownian motion  $\{B_t\}_{t \in [0, \infty)}$  on some prob. space  $(\Omega, \mathcal{A}, \mathbb{P})$ .  
 By disregarding a set of measure 0, we may assume that  $t \mapsto B_t(\omega)$  is cont. for every  $\omega \in \Omega$ .

Define  $B: \Omega \rightarrow \mathbb{X} = C([0, \infty))$  by  $\omega \mapsto (t \in [0, \infty) \mapsto B_t(\omega))$

Then for each  $t \in [0, \infty)$  we have the commutative diagram:



Since  $B_t$  is measurable for each  $t \in [0, \infty)$ , the map is measurable. Hence  $W := B_t(\mathbb{P})$

is a Borel prob. measure on  $\mathbb{X}$ , and if  $F \subseteq [0, \infty)$  is finite, then

$$\begin{aligned} \textcircled{79} \quad \mu_F &= (\pi_F)_* (W) = (\pi_F)_* (B_* (\mathbb{P})) \\ &= (\pi_F \circ B) (\mathbb{P}) = (B_F)_* (\mathbb{P}), \end{aligned}$$

where  $B_F(\omega) := (B_{t_1}(\omega), \dots, B_{t_n}(\omega))$ .

Hence  $\{\pi_t\}_{t \in [0, \infty)}$  is a Brownian motion obtained on  $\mathbb{R}^n$ .  $\square$

### §.24 Brownian motion on $\mathbb{R}^n$

An  $\mathbb{R}^n$ -valued stochastic process  $\{B_t\}_{t \in [0, \infty)}$  is called a (version of) Brownian motion on  $\mathbb{R}^n$  if the following conditions are true:

(i) the process is an  $\mathbb{R}^n$ -valued Gaussian process, i.e., for all  $k \in \mathbb{N}$ ,  $t_1 < \dots < t_k$ , the  $\mathbb{R}^{nk}$ -valued random variable  $(B_{t_1}, \dots, B_{t_k})$  has a Gaussian distribution.

Let

$$B_t^i = (B_t^1, \dots, B_t^n), \text{ where } B_t^i \text{ is } \mathbb{R}^1\text{-valued.}$$

(ii)  $T_{i, \infty} B_t^i$  is centered for  $i \in \{1, \dots, n\}$ ,  $t \in [0, \infty)$   
 $E[B_t^i] = 0$

(iii)  $\text{Cov}(B_s^i, B_t^j) = \delta_{ij} \wedge t$ ,  $i, j \in \{1, \dots, n\}$ ,  $s, t \in [0, \infty)$ .

(iv) sample paths  $t \mapsto B_t$  are cont. a.s.

(80) Rem. 8.25 (i) If  $B_t = (B_t^1, \dots, B_t^n)$  is a Brownian motion on  $\mathbb{R}^n$ , then  $B_t^1, \dots, B_t^n$  are ind. Brownian motions on  $\mathbb{R}$ .  
 Conversely, if  $B_t^1, \dots, B_t^n$  are ind. Brownian motions on  $\mathbb{R}$ , then  $B_t = (B_t^1, \dots, B_t^n)$  is a Brownian motion on  $\mathbb{R}^n$ .

(This proves existence!)

(ii) Uniqueness: One can show (as in Thm. 8.23) that there is a unique Wiener measure  $\mathbb{P} = \mathbb{C}([0, \infty), \mathbb{R}^n)$   
 $= \{ f: [0, \infty) \rightarrow \mathbb{R}^n \text{ cont.} \}$   
 s.t.  $\{ \pi_t \}_{t \in [0, \infty)}$  is a Brownian motion, where  $\pi_t: \mathbb{X} \rightarrow \mathbb{R}^n, f \mapsto f(t)$ .

Described by "marginals" on  $\mathbb{R}^n$   
 $\mathbb{P}_F := (\pi_F)_t(w) \quad \pi_F: \mathbb{X} \rightarrow (\mathbb{R}^n)^F$   
 $F \in [0, \infty)$  finite.  $f \mapsto f|_F$

(i)  $B_t = (B_t^1, \dots, B_t^n), t \in [0, \infty)$  is an  $\mathbb{R}^n$ -valued Brownian motion iff  
 $W_t := \lambda_1 B_t^1 + \dots + \lambda_n B_t^n$  is a 1-dim. Brownian motion for each unit vector  $(\lambda_1, \dots, \lambda_n) \in \mathbb{R}^n$ .  
 $\rightarrow$  Clear:  $\text{Cov}(W_s, W_t) = \lambda_1^2 s \wedge t + \dots + \lambda_n^2 s \wedge t = s \wedge t$

§1  $\leftarrow$  Need Sect:

$Z_1, \dots, Z_n$   $\mathbb{R}^k$ -valued r.v.'s.

Then they have a joint Gaussian distribution iff

$\lambda_1 Z_1 + \dots + \lambda_n Z_n$  is Gaussian for all  $\lambda_1, \dots, \lambda_n \in \mathbb{R}$ .

Details left as exercise!

## §.26. Basic properties of Brownian motion

Let  $\{B_t\}_{t \in [0, \infty)}$  be a Brownian motion in  $\mathbb{R}^n$ . Then the following processes are also Brownian motions:

(i)  $W_t := B_{t+s} - B_s$ ,  $s \in [0, \infty)$  fixed.

(Markov property)

Brownian motion is memoryless!

(ii) if  $A$  is an orthogonal transformation, then

$$W_t := A B_t$$

(iii)  $W_t = \frac{1}{a} B_{a^2 t}$ ,  $a > 0$  fixed,

(Brownian scaling)

(iv)  $W_t = \begin{cases} B_0 & t=0 \\ t B_{1/t} & t > 0 \end{cases}$

(time inversion.)

82 Proof: All processes  $W_t$  in (i) - (iv) are Gaussian, and  $W_t$  is centered.

One checks covariance: for example in (iii) and (iv):

$$\text{Cov}(W_s^k, W_t^l) = \text{Cov}\left(\frac{1}{a} B_{as}^k, \frac{1}{a} B_{at}^l\right) \\ = \frac{1}{a^2} \int_{k,l} (a^2 s \wedge a^2 t) = \int_{k,l} (s \wedge t)$$

$$\text{Cov}(W_s^k, W_t^l) = s \wedge t \text{ Cov}(B_{1/s}^k, B_{1/t}^l) \\ s, t > 0 \\ = s \wedge t \int_{k,l} \frac{1}{s} \wedge \frac{1}{t} \\ = \int_{k,l} t \wedge s, \quad \checkmark$$

Almost sure continuity of sample paths clear for: (i) - (iii), and on  $(0, \infty)$  in (iv): (ap t. measure 0)

Continuity at  $W_t$  at 0 is the following event:

$$A = \bigcap_{\varepsilon > 0} \bigcup_{\delta > 0} \bigcap_{t \in (0, \delta)} \left\{ \omega : |W_t(\omega) - W_0(\omega)| < \varepsilon \right\} \\ \varepsilon \in \mathbb{Q} \quad \delta \in \mathbb{Q} \quad t \in \mathbb{Q}$$

If we replace  $W_t$  by  $B_{1/t}$  then this is an almost sure event: since  $W_t$  and  $B_{1/t}$  have the same marginals, it follows that  $A$  is almost sure!

(Note: this shows that

$$\lim_{s \rightarrow \infty} \frac{|B_s|}{s} = 0 \quad \text{a.s.})$$

83) 8.27. The stochastic Loewner equation (SLE)

Chordal Loewner equation:

$\{f_t\}_{t \in [0, \infty)}$  normalized chordal Loewner chain

$$f_t(z) = z - \frac{2t}{z} + \dots \quad \text{near } \infty$$



Loewner-Kufner eq.

$$\frac{\partial f_t}{\partial t}(z, t) = V(z, t) \frac{\partial f_t}{\partial z}(z, t)$$

$$V(z, t) = 2 \int_{\mathbb{R}} \frac{d\nu_t(u)}{u-z}$$

$\nu_t$  prob. measure,  $\text{supp}(\nu_t) \subset \mathbb{R}$ .

One obtains  $SLE_\kappa$   $\kappa \geq 0$  if one takes a probabilistic driving function

$$\nu_t = \int_{\mathbb{R}} \sqrt{\kappa} B_t(u) \quad , \quad \text{where } B_t \text{ is 1-dim. Brownian motion.}$$

Then

$$V(z, t) = \frac{2}{\sqrt{\kappa} B_t(\omega) - z}$$

Depending on  $\omega$ , one gets a "harder"

84) Loewner chains and conv. random  
walks  $A_{\pm}(w)$

One is interested in these walks,  
because they can be used to study  
many conformally invariant processes  
in the plane.

Problems: 1) What are the characterizing  
properties of SLE?

(i.i.d. increments, Markov (= memoryless)  
property, conformal invariance)

2) What are the techniques to study  
SLE?

(Martingale methods, etc.)

## 9. Survey of martingale theory

### 9.1. Conditional expectation

Ex. Random exp. in two stages:  $\{1, \dots, 6\}$   
roll two dice with outcomes  $\underline{X}_1, \underline{X}_2$ ,  
 $Z = \underline{X}_1 + \underline{X}_2$ .  $\Omega = \{1, \dots, 6\} \times \{1, \dots, 6\}$   
 $E[Z] = 7$ .

Suppose outcome of  $\underline{X}_1$  is known (partial  
info.)  
Then we have to adjust  
 $E[Z]$  depending on  $\underline{X}_1$ .

$$E[Z | \underline{X}_1 = x] = x + 3.5 = \underline{X}_1(\omega) + 3.5.$$

New random vari.  $\Omega \ni \omega \mapsto E[Z | \underline{X}_1]$ .

85) Thm. and Def. 9.2. (conditional  
 $(\Omega, \mathcal{A}, \mathbb{P})$  prob. space, expectation)  
 $X$  r.v. with  $E[|X|] < \infty$ ,  $\mathcal{B} \subseteq \mathcal{A}$   
 $\mathcal{B}$ -algebra. Then there ex. a r.v.  $Y$   
 on  $\Omega$  s.t.

(i)  $Y$  is  $\mathcal{B}$ -measurable,

(ii)  $E[|Y|] < \infty$

(iii) for every  $B' \in \mathcal{B}$  we have

$$E[Y; B] = \int_{B'} Y(\omega) d\mathbb{P}(\omega)$$

$$= \int_{B'} X(\omega) d\mathbb{P}(\omega) = E[X; B]$$

$Y$  is essentially uniquely determined: if  
 $\hat{Y}$  is another r.v. with prop. (i)-(iii),  
 then  $Y = \hat{Y}$  a.s.

The r.v.  $Y$  is called (a version of)  
 conditional expectation of  $X$  given  $\mathcal{B}$ ,  
 denoted  $E[X | \mathcal{B}]$ .

Idea of proof: Wlog  $X \geq 0$ .

Define

$$\mu(B) := \int_B X(\omega) d\mathbb{P}(\omega) \quad \text{for } B \in \mathcal{B}.$$

Then  $\mu \ll \mathbb{P} | \mathcal{B}$ , so  $\mu$  has a Radon-  
 Nikodye derivative  $Y$  w.r.t.  $\mathbb{P} | \mathcal{B}$ .

Then (i) - (iii) are evident.

Uniqueness also clear.

$$E[X | Z_1, Z_2, \dots, Z_n] = E[X | \sigma(Z_1, \dots, Z_n)]$$

86) 9.3. Properties of conditional expectation

$(\Omega, \mathcal{A}, \mathbb{P})$  prob. space, all r.v.'s  $X$  satisfy  $E[|X|] < \infty$ ,

$\mathcal{B} \subseteq \mathcal{A}$   $\sigma$ -algebra.

- (i) If  $Y = E[X | \mathcal{B}]$ , then  $E[Y] = E[X]$ .
- (ii) If  $X$  is  $\mathcal{B}$ -measurable, then  $E[X | \mathcal{B}] = X$  (a.s.)
- (iii) Linearity
- (iv) if  $X \geq 0$ , then  $E[X | \mathcal{B}] \geq 0$
- (v)  $X_n \geq 0$ ,  $X_n \nearrow X$  (monotone a.s. convergence)

Then

$$E[X_n | \mathcal{B}] \nearrow E[X | \mathcal{B}] \text{ a.s.}$$

(Dominated conv.)

+ Fatou's lemma

$$|X_n(\omega)| \leq V(\omega), \quad E[V] < \infty,$$

$X_n \rightarrow X$  a.s. Then

$$E[X_n | \mathcal{B}] \rightarrow E[X | \mathcal{B}] \text{ a.s.}$$

(vii) (Jensen)  $\varphi: \mathbb{R} \rightarrow \mathbb{R}$  convex,

$$E[|\varphi(X)|] < \infty, \text{ then}$$

$$\varphi(E[X | \mathcal{B}]) \leq E[\varphi(X) | \mathcal{B}] \text{ a.s.}$$

In particular:  $\|E[X | \mathcal{B}]\| \leq \|X\|$   
and

(vii) (Tower property)  $\mathcal{E} \subseteq \mathcal{B} \subseteq \mathcal{A}$   $\sigma$ -algebras. Then

$$E[E[X | \mathcal{B}] | \mathcal{E}] = E[X | \mathcal{E}].$$

$$E[X | \mathcal{B} | \mathcal{E}]$$

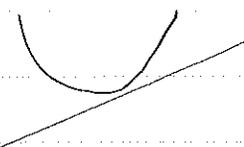
(87) (ix) If  $Z$  is  $\mathcal{B}$ -measurable, then  

$$E[Z X | \mathcal{B}] = Z E[X | \mathcal{B}]$$

(x) If  $X$  and  $\mathcal{B}$  are ind., then  

$$E[X | \mathcal{B}] = \underbrace{E[X]}_{\text{const. valued.}} \quad \text{a.s.}$$

Proof: Mostly straight forward from the definitions:

(vii) Jensen:   $\varphi$  convex

$$\varphi(x) = \sup_{L \in \mathcal{C} \text{ affine}} L(x)$$

$$L \in \mathcal{C} : L(X) = aX + b \leq \varphi(X)$$

$$\begin{aligned} \text{s.o.} \quad L(E[X | \mathcal{B}]) &\stackrel{\text{linearity}}{\leq} E[L(X) | \mathcal{B}] \\ &\stackrel{\text{monotonicity}}{\leq} E[\varphi(X) | \mathcal{B}] \end{aligned}$$

Taking sup over all  $L$  gives

$$\varphi(E[X | \mathcal{B}]) \leq E[\varphi(X) | \mathcal{B}]$$

Incorrect proof, because we take a sup over an uncountable family; can be corrected if we write

$$\varphi = \sup_{L_n \in \mathcal{C}} L_n \quad \text{for a countable collection } L_n, n \in \mathbb{N}.$$

(vi) For dominated convergence we need

Fatou Lemma  Jensen

$$\circ \neq X_n: \quad \text{Then } \quad \begin{aligned} &E[\liminf_{n \rightarrow \infty} X_n | \mathcal{B}] \stackrel{\text{Fatou}}{\leq} \liminf_{n \rightarrow \infty} E[X_n | \mathcal{B}] \\ &E[\liminf_{n \rightarrow \infty} X_n | \mathcal{B}] \stackrel{\text{Jensen}}{\geq} \liminf_{n \rightarrow \infty} E[X_n | \mathcal{B}] \end{aligned}$$

ⓂⓂ Ex. 9.4.  $(\Omega, \mathcal{A}, \mathbb{P})$  prob. space,  
 $\{A_n\}_{n \in \mathbb{N}}$  countable partition of  $\Omega$   
 with  $A_n \in \mathcal{A}$ ,  $\mathbb{P}(A_n) > 0$ .

$$\mathcal{B} = \sigma(\{A_n\}_{n \in \mathbb{N}}) = \left\{ \bigcup_{S \subseteq \mathbb{N}} A_n : S \text{ countable} \right\}.$$

Then

$$\mathbb{E}[X | \mathcal{B}] = \sum_{n \in \mathbb{N}} \frac{1}{\mathbb{P}(A_n)} \int_{A_n} X(\omega) d\mathbb{P}(\omega) \cdot \chi_{A_n}(\omega)$$

$$= \sum_{n \in \mathbb{N}} \mathbb{E}[X; A_n] \cdot \chi_{A_n}$$

Check definition!

Ex. 9.5. (Fair games and martingales)

Players I + II roll dice.

Zero-sum game, at each step

P1 wins or loses 1 unit.

$X_n$  winnings of P1 after  $n$  rolls.

$(-X_n)$  " " P2 after  $n$  rolls.)

Game 1: P1 wins if roll  $\in \{1, 2\}$  not fair!  
 (s. loses if  $\in \{3, 4, 5, 6\}$ )

G2: P1 wins if roll even fair!

G3: P1 wins if roll even; fair!  
 if <sup>and</sup> the player's loss  $\geq 100$  units,  
 then game biased against player  
 as in G1.

G3 is a fair game ( $\mathbb{E}[X_n] = 0$  for all  $n \in \mathbb{N}$ ),

but not fair at all times (or all situations).

89) How to model a game that is "fair at all times".

$$\mathbb{E}[X_{n+1} - X_n] = 0 \quad (\text{true if } \mathbb{E}[X_{n+1}] = \mathbb{E}[X_n] = 0)$$

Better:  $\mathbb{E}[X_{n+1} - X_n \mid \mathcal{F}_n = x] = 0$   
whatever  $x$ .

$\mathcal{F}_n$   $\sigma$ -algebra of events that will be known by time  $n$ .

$$\mathbb{E}[X_{n+1} - X_n \mid \mathcal{F}_n] = 0 \quad \text{equiv.}$$

$$\mathbb{E}[X_{n+1} \mid \mathcal{F}_n] = X_n \quad (\mathbb{E}[X_n \mid \mathcal{F}_n] = X_n)$$

Def. 9.6: (Martingales; discrete-time case).

$(\Omega, \mathcal{A}, \mathbb{P})$  prop. space

with filtration given  $\sigma$ -algebras

$$\mathcal{F}_n \subseteq \mathcal{A} \quad \text{for } n \in \mathbb{N}_0 \quad (\mathcal{F}_n \subseteq \mathcal{F}_{n+1} \text{ for } n \in \mathbb{N}_0)$$

$X = \{X_n\}_{n \in \mathbb{N}_0}$  sequence of r.v.'s on  $\Omega$ .

Then  $X$  is called a martingale if

(i)  $X_n$  is  $\mathcal{F}_n$ -measurable for  $n \in \mathbb{N}_0$ ,  
and  $X_n \in L^1$ , i.e.,  $\mathbb{E}[|X_n|] < \infty$ .

(ii)  $\mathbb{E}[X_{n+1} \mid \mathcal{F}_n] = X_n$  (a.s.) for  $n \in \mathbb{N}_0$ .

If in (ii) we have  $\leq$  or  $\geq$ , then

$X$  is called a super-martingale

or submartingale, resp.

submartingale: tendency to increase

supermartingale: tendency to decrease

⑨ Often:  $\mathcal{F}_n = \sigma(X_0, \dots, X_n), n \in \mathbb{N}$ .  
 "natural filtration"

Ex. 9.7a) Games as in 9.5, with natural filtration,  $\mathcal{X} = \{X_n\}_{n \in \mathbb{N}}$ .

$G_1, G_3$  not martingales,  $G_2, G_4$  martingales.

$G_1$  supermartingale

b) (Dyadic martingales)

$\Omega = [0, 1]$  with Lebesgue measure,  
 $f \in L^1[0, 1]$ .

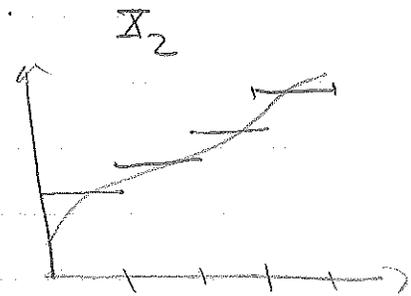
$D_{n,k} = [\frac{k}{2^n}, \frac{k+1}{2^n}]$ ,  $n \in \mathbb{N}, k = 1, \dots, 2^n - 1$   
 dyadic interval

$\mathcal{F}_n$   $\sigma$ -algebra generated by dyadic intervals at level  $\leq n$ .

$$X_n(\omega) = \sum_{k=1}^{2^n-1} \chi_{D_{n,k}}(\omega) \int_{D_{n,k}} f$$

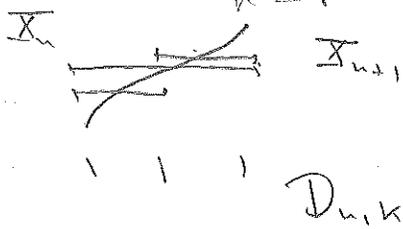
$n \in \mathbb{N}$ .

$\mathcal{X} = \{X_n\}_{n \in \mathbb{N}}$  is martingale.



(i)  $X_n$  is  $\mathcal{F}_n$ -measurable,

(ii) 
$$\mathbb{E}[X_{n+1} | \mathcal{F}_n] = \sum_{k=1}^{2^n-1} \chi_{D_{n,k}}(\omega) \int_{D_{n,k}} X_{n+1} = X_n.$$



Note:  $X_n(\omega) \rightarrow f(\omega)$

for a.e.  $\omega$ .

Instance of martingale conv. thm!

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c) (Brownian motion)  
 $0 \leq t_0 < t_1 < \dots$   $B_t$  Brownian motion on  $\mathbb{R}$ .

$X_n = B_{t_n}$  for  $n \in \mathbb{N}_0$ .  
Then  $X = \{X_n\}_{n \in \mathbb{N}_0}$  (with natural filtration)

is a martingale.

$B_{t_{n+1}} - B_{t_n}$  ind. of  $B_{t_0}, \dots, B_{t_n}$

$$E[B_t] = 0.$$

(i)  $X_n = B_{t_n}$  is  $\mathcal{F}_n = \sigma(B_{t_0}, \dots, B_{t_n})$

$$(ii) E[X_{n+1} | \mathcal{F}_n] = E[B_{t_{n+1}} | \mathcal{F}_n]$$

$$= E[B_{t_{n+1}} - B_{t_n} | \mathcal{F}_n] + B_{t_n}$$

$$= E[B_{t_{n+1}} - B_{t_n}] + B_{t_n}$$

$$= B_{t_n} = X_n$$

Def. 9.8. (Martingales; continuous-time case).

$(\Omega, \mathcal{A}, \mathbb{P})$  prob. space

$\{\mathcal{F}_t\}_{t \geq 0}$  filtration (i.e.  $\mathcal{F}_t \subseteq \mathcal{A}$ )

and  $\mathcal{F}_s \subseteq \mathcal{F}_t$  for  $s \leq t$   $\mathcal{G}$ -algebra

A stochastic process (after extra technical conditions)

is called adapted if  $X_t$  is  $\mathcal{F}_t$ -measurable

92) For all  $t \geq 0$ .

$\underline{X} = \{X_t\}_{t \geq 0}$  is a martingale if

(i)  $\underline{X}$  is adapted and  $E[|X_t|] < \infty$   
for  $t \geq 0$ ,

(iii)  $E[X_t | \mathcal{F}_s] = X_s$  for all  
 $0 \leq s \leq t$ .

Natural filtration  $\mathcal{F}_t = \sigma(X_s : 0 \leq s \leq t)$ .

Ex. 9.9 a)  $\{B_t\}$  Brownian motion (with  
natural filtration is martingale.

b)  $B_t$  Brownian motion  
 $\mathcal{F}_t = \sigma(B_s : 0 \leq s \leq t)$ .

$\underline{X}_t = B_t^2 - t$  is a martingale.

(i)  $\{X_t\}$  is adapted, and  $E[|X_t|] < \infty$ .

(ii)  $E[X_t | \mathcal{F}_s] = E[B_t^2 - t | \mathcal{F}_s]$   
 $s \leq t$

$$= E[(B_t - B_s + B_s)^2 | \mathcal{F}_s] - t$$

$$= E[(B_t - B_s)^2 + 2B_s(B_t - B_s) + B_s^2 | \mathcal{F}_s] - t$$

$$= E[(B_t - B_s)^2 | \mathcal{F}_s] + 2B_s E[B_t - B_s | \mathcal{F}_s]$$

$$+ B_s^2 - t = E[(B_t - B_s)^2] + 2B_s E[B_t - B_s]$$

$$+ B_s^2 - t = (t-s) + B_s^2 - t = B_s^2 - s = X_s.$$

93) Converse:

Thm. (Lévy)  $\{X_t\}_{t \geq 0}$  continuous martingale (i.e., martingale with almost surely cont. sample paths).

If  $\{X_t^2 - t\}_{t \geq 0}$  is a martingale (w.r.t.  $\mathcal{F}_t = \sigma(X_s, s \leq t)$ ),

then  $\{X_t\}_{t \geq 0}$  is a Brownian motion.

Important facts about martingales:  
martingale convergence thm.,  
Doob's  $L^p$ -submartingale inequalities,  
sub- and supermartingale decompositions,  
optional stopping, stochastic integrals.